

Marginalization, Conditioning, and Bayes Rule

Prof. James Richard Forbes

McGill University, Department of Mechanical Engineering



The Gaussian Distribution

- ▶ A continuous random variable $\mathbf{x} \in \mathbb{R}^{n_x}$ is said to have a normal or Gaussian distribution if the pdf associated with $\mathbf{x}$ is given by

$$p(\mathbf{x}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^{n_x} \det \boldsymbol{\Sigma}}} \exp \left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right), \quad (1)$$

where $\bar{\mathbf{x}}$ is the mean and $\boldsymbol{\Sigma} = \boldsymbol{\Sigma}^\top > 0$ is the covariance matrix.

- ▶ $\boldsymbol{\Sigma} = \boldsymbol{\Sigma}^\top > 0$ means $\boldsymbol{\Sigma}$ is symmetric and positive definite, thus ensuring $\boldsymbol{\Sigma}$ is not singular, and thus $\boldsymbol{\Sigma}^{-1}$ exists.
- ▶ Often it is more convenient to write the Gaussian distribution in information form,

$$p(\mathbf{x}) = \mathcal{N}^{-1}(\boldsymbol{\eta}, \boldsymbol{\Lambda}) = \mathcal{N}^{-1}(\mathbf{x}; \boldsymbol{\eta}, \boldsymbol{\Lambda}) = \frac{\exp(-\frac{1}{2} \boldsymbol{\eta}^\top \boldsymbol{\Lambda}^{-1} \boldsymbol{\eta})}{\sqrt{(2\pi)^{n_x} \det \boldsymbol{\Lambda}^{-1}}} \exp \left(-\frac{1}{2} \mathbf{x}^\top \boldsymbol{\Lambda} \mathbf{x} + \mathbf{x}^\top \boldsymbol{\eta} \right), \quad (2)$$

where $\boldsymbol{\Lambda} = \boldsymbol{\Sigma}^{-1}$, and $\boldsymbol{\eta} = \boldsymbol{\Lambda} \boldsymbol{\mu}$.

- Equation (2) can be derived from (1). Starting from (1),

$$\begin{aligned}\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) &= \frac{1}{\sqrt{(2\pi)^{n_x} \det \boldsymbol{\Sigma}}} \exp \left(-\frac{1}{2} \left(\mathbf{x}^\top \boldsymbol{\Sigma}^{-1} \mathbf{x} - 2\mathbf{x}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} + \boldsymbol{\mu}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right) \right) \\&= \frac{\exp \left(\boldsymbol{\mu}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right)}{\sqrt{(2\pi)^{n_x} \det \boldsymbol{\Sigma}}} \exp \left(-\frac{1}{2} \mathbf{x}^\top \boldsymbol{\Sigma}^{-1} \mathbf{x} + \mathbf{x}^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right) \\&= \frac{\exp \left(-\frac{1}{2} \boldsymbol{\eta}^\top \boldsymbol{\Lambda}^{-1} \boldsymbol{\eta} \right)}{\sqrt{(2\pi)^{n_x} \det \boldsymbol{\Lambda}^{-1}}} \exp \left(-\frac{1}{2} \mathbf{x}^\top \boldsymbol{\Lambda} \mathbf{x} + \mathbf{x}^\top \boldsymbol{\eta} \right) \\&= \mathcal{N}^{-1}(\boldsymbol{\eta}, \boldsymbol{\Lambda}).\end{aligned}\tag{3}$$

- A short-hand notation for indicating $\mathbf{x}$ is normally distributed is

$$\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathcal{N}^{-1}(\boldsymbol{\eta}, \boldsymbol{\Lambda}) = \mathcal{N}^{-1}(\mathbf{x}; \boldsymbol{\eta}, \boldsymbol{\Lambda}).$$

Three Critical Operations

- Marginalization

$$p(\mathbf{x}) = \int_{-\infty}^{\infty} p(\mathbf{x}, \mathbf{y}) d\mathbf{y} \quad (4)$$

- Conditioning

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{x}|\mathbf{y})p(\mathbf{y}) \quad (5)$$

- Bayes Rule

$$p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{x})p(\mathbf{x})}{p(\mathbf{y})} \quad (6)$$

Three Critical Operations (Conditional Form)

► Marginalization

$$p(\mathbf{x}|\boldsymbol{\theta}) = \int_{-\infty}^{\infty} p(\mathbf{x}, \mathbf{y}|\boldsymbol{\theta}) d\mathbf{y} \quad (7)$$

► Conditioning

$$p(\mathbf{x}, \mathbf{y}|\boldsymbol{\theta}) = p(\mathbf{x}|\mathbf{y}, \boldsymbol{\theta})p(\mathbf{y}|\boldsymbol{\theta}) \quad (8)$$

► Bayes Rule

$$p(\mathbf{x}|\mathbf{y}, \boldsymbol{\theta}) = \frac{p(\mathbf{y}|\mathbf{x}, \boldsymbol{\theta})p(\mathbf{x}|\boldsymbol{\theta})}{p(\mathbf{y}|\boldsymbol{\theta})} \quad (9)$$

The Schur Complement

Theorem (1)

[1] Consider the matrix $\mathbf{X} \in \mathbb{R}^{n \times n}$. Partition $\mathbf{X}$ in block matrix form

$$\mathbf{X} = \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix}, \quad (10)$$

where $\mathbf{A} \in \mathbb{R}^{p \times p}$, $\mathbf{B} \in \mathbb{R}^{p \times q}$, $\mathbf{C} \in \mathbb{R}^{q \times p}$, and $\mathbf{D} \in \mathbb{R}^{q \times q}$. If both $\mathbf{D}$ and $\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C}$ are invertible then

$$\mathbf{X}^{-1} = \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix}^{-1} = \begin{bmatrix} (\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1} & -(\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1}\mathbf{BD}^{-1} \\ -\mathbf{D}^{-1}\mathbf{C}(\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1} & \mathbf{D}^{-1} + \mathbf{D}^{-1}\mathbf{C}(\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1}\mathbf{BD}^{-1} \end{bmatrix} \quad (11)$$

The matrix $\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C}$ is the Schur Complement of $\mathbf{D}$ in $\mathbf{X}$.

Proof.

- Decompose $\mathbf{X}$ into three matrices,

$$\mathbf{X} = \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix} = \underbrace{\begin{bmatrix} \mathbf{1} & \mathbf{BD}^{-1} \\ \mathbf{0} & \mathbf{1} \end{bmatrix}}_{\mathbf{P}} \underbrace{\begin{bmatrix} \mathbf{A} - \mathbf{BD}^{-1}\mathbf{C} & \mathbf{0} \\ \mathbf{0} & \mathbf{D} \end{bmatrix}}_{\mathbf{Q}} \underbrace{\begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \mathbf{D}^{-1}\mathbf{C} & \mathbf{1} \end{bmatrix}}_{\mathbf{R}} \quad (12)$$

- Note that

$$\begin{aligned} \mathbf{P}^{-1} &= \begin{bmatrix} \mathbf{1} & \mathbf{BD}^{-1} \\ \mathbf{0} & \mathbf{1} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{1} & -\mathbf{BD}^{-1} \\ \mathbf{0} & \mathbf{1} \end{bmatrix}, \\ \mathbf{Q}^{-1} &= \begin{bmatrix} \mathbf{A} - \mathbf{BD}^{-1}\mathbf{C} & \mathbf{0} \\ \mathbf{0} & \mathbf{D} \end{bmatrix}^{-1} = \begin{bmatrix} (\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{D}^{-1} \end{bmatrix}, \\ \mathbf{R}^{-1} &= \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \mathbf{D}^{-1}\mathbf{C} & \mathbf{1} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ -\mathbf{D}^{-1}\mathbf{C} & \mathbf{1} \end{bmatrix}. \end{aligned}$$

- Recall that, for any three invertible matrices $\mathbf{P}, \mathbf{Q}, \mathbf{R} \in \mathbb{R}^{n \times n}$,

$$(\mathbf{PQR})^{-1} = \mathbf{R}^{-1}\mathbf{Q}^{-1}\mathbf{P}^{-1}.$$

- Therefore,

$$\begin{aligned} \mathbf{X}^{-1} &= (\mathbf{PQR})^{-1} = \mathbf{R}^{-1}\mathbf{Q}^{-1}\mathbf{P}^{-1} = \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ -\mathbf{D}^{-1}\mathbf{C} & \mathbf{1} \end{bmatrix} \begin{bmatrix} (\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{D}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{1} & -\mathbf{BD}^{-1} \\ \mathbf{0} & \mathbf{1} \end{bmatrix} \\ &= \begin{bmatrix} (\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1} & -(\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1}\mathbf{BD}^{-1} \\ -\mathbf{D}^{-1}\mathbf{C}(\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1} & \mathbf{D}^{-1} + \mathbf{D}^{-1}\mathbf{C}(\mathbf{A} - \mathbf{BD}^{-1}\mathbf{C})^{-1}\mathbf{BD}^{-1} \end{bmatrix}. \end{aligned} \quad (11)$$

The Schur Complement (Alternative Version)

Theorem (2)

[1] Consider the matrix $\mathbf{X} \in \mathbb{R}^{n \times n}$. Partition $\mathbf{X}$ in block matrix form

$$\mathbf{X} = \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix}, \quad (13)$$

where $\mathbf{A} \in \mathbb{R}^{p \times p}$, $\mathbf{B} \in \mathbb{R}^{p \times q}$, $\mathbf{C} \in \mathbb{R}^{q \times p}$, and $\mathbf{D} \in \mathbb{R}^{q \times q}$. If both $\mathbf{A}$ and $\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B}$ are invertible then

$$\mathbf{X}^{-1} = \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{A}^{-1} + \mathbf{A}^{-1}\mathbf{B}(\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1}\mathbf{CA}^{-1} & -\mathbf{A}^{-1}\mathbf{B}(\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1} \\ -(\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1}\mathbf{CA}^{-1} & (\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1} \end{bmatrix} \quad (14)$$

The matrix $\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B}$ is the Schur Complement of $\mathbf{A}$ in $\mathbf{X}$.

Proof.

- Decompose $\mathbf{X}$ into three matrices,

$$\mathbf{X} = \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix} = \underbrace{\begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \mathbf{BA}^{-1} & \mathbf{1} \end{bmatrix}}_{\mathbf{P}} \underbrace{\begin{bmatrix} \mathbf{A} & \mathbf{0} \\ \mathbf{0} & \mathbf{D} - \mathbf{CA}^{-1}\mathbf{B} \end{bmatrix}}_{\mathbf{Q}} \underbrace{\begin{bmatrix} \mathbf{1} & \mathbf{A}^{-1}\mathbf{C} \\ \mathbf{0} & \mathbf{1} \end{bmatrix}}_{\mathbf{R}} \quad (15)$$

- Note that

$$\mathbf{P}^{-1} = \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \mathbf{BA}^{-1} & \mathbf{1} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ -\mathbf{BA}^{-1} & \mathbf{1} \end{bmatrix},$$

$$\mathbf{Q}^{-1} = \begin{bmatrix} \mathbf{A} & \mathbf{0} \\ \mathbf{0} & \mathbf{D} - \mathbf{CA}^{-1}\mathbf{B} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{A}^{-1} & \mathbf{0} \\ \mathbf{0} & (\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1} \end{bmatrix},$$

$$\mathbf{R}^{-1} = \begin{bmatrix} \mathbf{1} & \mathbf{A}^{-1}\mathbf{C} \\ \mathbf{0} & \mathbf{1} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{1} & -\mathbf{A}^{-1}\mathbf{C} \\ \mathbf{0} & \mathbf{1} \end{bmatrix}.$$

- Recall that, for any three invertible matrices $\mathbf{P}, \mathbf{Q}, \mathbf{R} \in \mathbb{R}^{n \times n}$,

$$(\mathbf{PQR})^{-1} = \mathbf{R}^{-1}\mathbf{Q}^{-1}\mathbf{P}^{-1}.$$

- Therefore,

$$\begin{aligned} \mathbf{X}^{-1} &= (\mathbf{PQR})^{-1} = \mathbf{R}^{-1}\mathbf{Q}^{-1}\mathbf{P}^{-1} = \begin{bmatrix} \mathbf{1} & -\mathbf{A}^{-1}\mathbf{C} \\ \mathbf{0} & \mathbf{1} \end{bmatrix} \begin{bmatrix} \mathbf{A}^{-1} & \mathbf{0} \\ \mathbf{0} & (\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ -\mathbf{BA}^{-1} & \mathbf{1} \end{bmatrix} \\ &= \begin{bmatrix} \mathbf{A}^{-1} + \mathbf{A}^{-1}\mathbf{B}(\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1}\mathbf{CA}^{-1} & -\mathbf{A}^{-1}\mathbf{B}(\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1} \\ -(\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1}\mathbf{CA}^{-1} & (\mathbf{D} - \mathbf{CA}^{-1}\mathbf{B})^{-1} \end{bmatrix}. \end{aligned} \quad (14)$$

The Matrix Inversion Lemma

- ▶ In the special case where $\mathbf{A}$, $\mathbf{D}$, the Schur complement of $\mathbf{A}$ in $\mathbf{X}$, $\mathbf{A} - \mathbf{B}\mathbf{D}^{-1}\mathbf{C}$, and the Schur complement of $\mathbf{D}$ in $\mathbf{X}$, $\mathbf{D} - \mathbf{C}\mathbf{A}^{-1}\mathbf{B}$, are all invertible, equating the $(1, 1)$ block of each expression for $\mathbf{X}^{-1}$ in Equations (11) and (14) leads to [1]

$$(\mathbf{A} - \mathbf{B}\mathbf{D}^{-1}\mathbf{C})^{-1} = \mathbf{A}^{-1} + \mathbf{A}^{-1}\mathbf{B}(\mathbf{D} - \mathbf{C}\mathbf{A}^{-1}\mathbf{B})^{-1}\mathbf{C}\mathbf{A}^{-1},$$

which is called the Woodbury matrix identity.

- ▶ The *Matrix Inversion Lemma* is,

$$(\mathbf{A} + \mathbf{B}\mathbf{C})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1}\mathbf{B}(\mathbf{1} + \mathbf{C}\mathbf{A}^{-1}\mathbf{B})^{-1}\mathbf{C}\mathbf{A}^{-1}.$$

which comes from setting $\mathbf{D} = \mathbf{1}$ and changing $\mathbf{B}$ to $-\mathbf{B}$ [1].

Determinant Results

Consider

$$\begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix},$$

where $\mathbf{A}$ and $\mathbf{D}$ are square.

- Assuming $\mathbf{A}$ is full rank,

$$\det \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix} = \det \mathbf{A} \det (\mathbf{D} - \mathbf{C}\mathbf{A}^{-1}\mathbf{B}), \quad (16)$$

which can be derived by taking the determinant of both sides of (15).

- Assuming $\mathbf{D}$ is full rank,

$$\det \begin{bmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{bmatrix} = \det \mathbf{D} \det (\mathbf{A} - \mathbf{B}\mathbf{D}^{-1}\mathbf{C}), \quad (17)$$

which can be derived by taking the determinant of both sides of (12).

Specialized Schur Complement Results

► Consider

$$\mathbf{z} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \quad \mathbf{z} = \begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}, \quad \boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xy} \\ \boldsymbol{\Sigma}_{yx} & \boldsymbol{\Sigma}_{yy} \end{bmatrix}, \quad \boldsymbol{\Lambda} = \begin{bmatrix} \boldsymbol{\Lambda}_{xx} & \boldsymbol{\Lambda}_{xy} \\ \boldsymbol{\Lambda}_{yx} & \boldsymbol{\Lambda}_{yy} \end{bmatrix}, \quad (18)$$

where $\boldsymbol{\Lambda} = \boldsymbol{\Sigma}^{-1}$, $\boldsymbol{\Sigma}_{yx} = \boldsymbol{\Sigma}_{xy}^\top$, and $\boldsymbol{\Lambda}_{yx} = \boldsymbol{\Lambda}_{xy}^\top$.

► The Schur complement of Σ_{yy} in Σ is

$$\begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Lambda_{xx} & \Lambda_{xy} \\ \Lambda_{yx} & \Lambda_{yy} \end{bmatrix} = \begin{bmatrix} (\Sigma_{xx} - \Sigma_{xy}\Sigma_{yy}^{-1}\Sigma_{yx})^{-1} & -(\Sigma_{xx} - \Sigma_{xy}\Sigma_{yy}^{-1}\Sigma_{yx})^{-1}\Sigma_{xy}\Sigma_{yy}^{-1} \\ -\Sigma_{yy}^{-1}\Sigma_{yx}(\Sigma_{xx} - \Sigma_{xy}\Sigma_{yy}^{-1}\Sigma_{yx})^{-1} & \Sigma_{yy}^{-1} + \Sigma_{yy}^{-1}\Sigma_{yx}(\Sigma_{xx} - \Sigma_{xy}\Sigma_{yy}^{-1}\Sigma_{yx})^{-1}\Sigma_{xy}\Sigma_{yy}^{-1} \end{bmatrix}. \quad (19)$$

► Thus,

$$\Lambda_{xx}^{-1} = \Sigma_{xx} - \Sigma_{xy}\Sigma_{yy}^{-1}\Sigma_{yx}. \quad (20)$$

► The Schur complement of Σ_{xx} in Σ is

$$\begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Lambda_{xx} & \Lambda_{xy} \\ \Lambda_{yx} & \Lambda_{yy} \end{bmatrix} = \begin{bmatrix} \Sigma_{xx}^{-1} + \Sigma_{xx}^{-1}\Sigma_{xy}(\Sigma_{yy} - \Sigma_{yx}\Sigma_{xx}^{-1}\Sigma_{xy})^{-1}\Sigma_{yx}\Sigma_{xx}^{-1} & -\Sigma_{xx}^{-1}\Sigma_{xy}(\Sigma_{yy} - \Sigma_{yx}\Sigma_{xx}^{-1}\Sigma_{xy})^{-1} \\ -(\Sigma_{yy} - \Sigma_{yx}\Sigma_{xx}^{-1}\Sigma_{xy})^{-1}\Sigma_{yx}\Sigma_{xx}^{-1} & (\Sigma_{yy} - \Sigma_{yx}\Sigma_{xx}^{-1}\Sigma_{xy})^{-1} \end{bmatrix}. \quad (21)$$

► Thus,

$$\Lambda_{yy}^{-1} = \Sigma_{yy} - \Sigma_{yx}\Sigma_{xx}^{-1}\Sigma_{xy}. \quad (22)$$

► The Schur complement of Λ_{yy} in Λ is

$$\begin{bmatrix} \Lambda_{xx} & \Lambda_{xy} \\ \Lambda_{yx} & \Lambda_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix} = \begin{bmatrix} (\Lambda_{xx} - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx})^{-1} & -(\Lambda_{xx} - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx})^{-1}\Lambda_{xy}\Lambda_{yy}^{-1} \\ -\Lambda_{yy}^{-1}\Lambda_{yx}(\Lambda_{xx} - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx})^{-1} & \Lambda_{yy}^{-1} + \Lambda_{yy}^{-1}\Lambda_{yx}(\Lambda_{xx} - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx})^{-1}\Lambda_{xy}\Lambda_{yy}^{-1} \end{bmatrix}. \quad (23)$$

► Thus,

$$\Sigma_{xx}^{-1} = \Lambda_{xx} - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx}. \quad (24)$$

► The Schur complement of Λ_{xx} in Λ is

$$\begin{bmatrix} \Lambda_{xx} & \Lambda_{xy} \\ \Lambda_{yx} & \Lambda_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix} = \begin{bmatrix} \Lambda_{xx}^{-1} + \Lambda_{xx}^{-1}\Lambda_{xy}(\Lambda_{yy} - \Lambda_{yx}\Lambda_{xx}^{-1}\Lambda_{xy})^{-1}\Lambda_{yx}\Lambda_{xx}^{-1} & -\Lambda_{xx}^{-1}\Lambda_{xy}(\Lambda_{yy} - \Lambda_{yx}\Lambda_{xx}^{-1}\Lambda_{xy})^{-1} \\ -(\Lambda_{yy} - \Lambda_{yx}\Lambda_{xx}^{-1}\Lambda_{xy})^{-1}\Lambda_{yx}\Lambda_{xx}^{-1} & (\Lambda_{yy} - \Lambda_{yx}\Lambda_{xx}^{-1}\Lambda_{xy})^{-1} \end{bmatrix}. \quad (25)$$

► Thus,

$$\Sigma_{yy}^{-1} = \Lambda_{yy} - \Lambda_{yx}\Lambda_{xx}^{-1}\Lambda_{xy}. \quad (26)$$

Marginalization Theorem

► Consider

$$\mathbf{z} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \quad \mathbf{z} = \begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}, \quad \boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xy} \\ \boldsymbol{\Sigma}_{yx} & \boldsymbol{\Sigma}_{yy} \end{bmatrix}, \quad \boldsymbol{\Lambda} = \begin{bmatrix} \boldsymbol{\Lambda}_{xx} & \boldsymbol{\Lambda}_{xy} \\ \boldsymbol{\Lambda}_{yx} & \boldsymbol{\Lambda}_{yy} \end{bmatrix}. \quad (18)$$

Theorem (3)

Let $\mathbf{z} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ and be partitioned as in (18). Then, the marginal pdf $p(\mathbf{x})$ is given by

$$\begin{aligned} p(\mathbf{x}) &= \mathcal{N}(\boldsymbol{\mu}_x, \boldsymbol{\Sigma}_{xx}) \\ &= \mathcal{N}^{-1}(\bar{\boldsymbol{\eta}}, \bar{\boldsymbol{\Lambda}}) \end{aligned}$$

where

$$\bar{\boldsymbol{\eta}} = \boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\eta}_y, \quad \bar{\boldsymbol{\Lambda}} = \boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}.$$

Proof.

► Recall that

$$p(\mathbf{x}) = \int_{-\infty}^{\infty} p(\mathbf{x}, \mathbf{y}) d\mathbf{y}. \quad (27)$$

► The joint pdf is

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^{(n_x+n_y)} \det \boldsymbol{\Sigma}}} \exp(\alpha) \quad (28)$$

where the argument of the exponent, α , is

$$\alpha = -\frac{1}{2} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right)^T \begin{bmatrix} \boldsymbol{\Lambda}_{xx} & \boldsymbol{\Lambda}_{xy} \\ \boldsymbol{\Lambda}_{yx} & \boldsymbol{\Lambda}_{yy} \end{bmatrix} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right). \quad (29)$$

► Using the Schur complement of $\boldsymbol{\Lambda}_{yy}$ in $\boldsymbol{\Lambda}$, Theorem 1, enables α to be written as

$$\begin{aligned} \alpha = -\frac{1}{2} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right)^T & \begin{bmatrix} \mathbf{1} & \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \\ \mathbf{0} & \mathbf{1} \end{bmatrix} \begin{bmatrix} \boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx} & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Lambda}_{yy} \end{bmatrix} \\ & \times \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx} & \mathbf{1} \end{bmatrix} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right). \end{aligned} \quad (30)$$

- ▶ Next, expand α , and write α as the sum of two quadratic terms:
 - ▶ a quadratic term in $(\mathbf{y} - \boldsymbol{\mu}_y) + \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}(\mathbf{x} - \boldsymbol{\mu}_x)$, called α_1 , and
 - ▶ a quadratic term in $\mathbf{x} - \boldsymbol{\mu}_x$, called α_2 .
- ▶ It follows that $\alpha = \alpha_1 + \alpha_2$ where

$$\alpha_1 = -\frac{1}{2} \left((\mathbf{y} - \boldsymbol{\mu}_y) + \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}(\mathbf{x} - \boldsymbol{\mu}_x) \right)^\top \boldsymbol{\Lambda}_{yy} \left((\mathbf{y} - \boldsymbol{\mu}_y) + \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}(\mathbf{x} - \boldsymbol{\mu}_x) \right), \quad (31)$$

$$\begin{aligned} \alpha_2 &= -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top (\boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}) (\mathbf{x} - \boldsymbol{\mu}_x) \\ &= -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \boldsymbol{\Sigma}_{xx}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x), \end{aligned} \quad (32)$$

and (24) has been used.

- ▶ The joint pdf $p(\mathbf{x}, \mathbf{y})$ can now written as

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^{(n_x+n_y)} \det \boldsymbol{\Sigma}}} \exp(\alpha_1) \exp(\alpha_2) \quad (33)$$

► Next,

$$\begin{aligned} p(\mathbf{x}) &= \int_{-\infty}^{\infty} p(\mathbf{x}, \mathbf{y}) d\mathbf{y} \\ &= \frac{1}{\sqrt{(2\pi)^{(n_x+n_y)} \det \Sigma}} \exp(\alpha_2) \int_{-\infty}^{\infty} \exp(\alpha_1) d\mathbf{y}, \end{aligned} \quad (34)$$

because $\exp(\alpha_2)$ does not depend on $\mathbf{y}$.

► The integral of $\exp(\alpha_1)$ over $\mathbf{y}$ is

$$\int_{-\infty}^{\infty} \exp(\alpha_1) d\mathbf{y} = \sqrt{(2\pi)^{n_y} \det \Lambda_{yy}^{-1}}, \quad (35)$$

which follows from the fact that

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{(2\pi)^{n_y} \det \Lambda_{yy}^{-1}}} \exp(\alpha_1) d\mathbf{y} = 1 \quad (36)$$

to satisfy the axiom of total probability.

- Therefore, the density $p(\mathbf{x})$ can be written as

$$p(\mathbf{x}) = \frac{\sqrt{\det \Lambda_{yy}^{-1}}}{\sqrt{(2\pi)^{n_x} \det \Sigma}} \exp \left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \Sigma_{xx}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x) \right). \quad (37)$$

- Using (22),

$$\Lambda_{yy}^{-1} = \Sigma_{yy} - \Sigma_{yx} \Sigma_{xx}^{-1} \Sigma_{xy} \quad (22)$$

and (16),

$$\det \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix} = \det \Sigma_{xx} \det(\Sigma_{yy} - \Sigma_{yx} \Sigma_{xx}^{-1} \Sigma_{xy}), \quad (38)$$

it follows that

$$p(\mathbf{x}) = \mathcal{N}(\boldsymbol{\mu}_x, \Sigma_{xx}) = \frac{1}{\sqrt{(2\pi)^{n_x} \det \Sigma_{xx}}} \exp \left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \Sigma_{xx}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x) \right). \quad (39)$$

- To derive the information form, first recall that

$$\begin{bmatrix} \eta_x \\ \eta_y \end{bmatrix} = \begin{bmatrix} \Lambda_{xx} & \Lambda_{xy} \\ \Lambda_{yx} & \Lambda_{yy} \end{bmatrix} \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix} = \begin{bmatrix} \Lambda_{xx}\mu_x + \Lambda_{xy}\mu_y \\ \Lambda_{yx}\mu_x + \Lambda_{yy}\mu_y \end{bmatrix},$$

or

$$\Lambda_{xx}\mu_x = \eta_x - \Lambda_{xy}\mu_y \quad (40)$$

$$\Lambda_{yx}\mu_x = \eta_y - \Lambda_{yy}\mu_y. \quad (41)$$

- Next, write the exponent α_2 as

$$\begin{aligned} \alpha_2 &= -\frac{1}{2}(\mathbf{x} - \mu_x)^\top \Sigma_{xx}^{-1}(\mathbf{x} - \mu_x) \\ &= -\frac{1}{2}\mathbf{x}^\top \Sigma_{xx}^{-1}\mathbf{x} + \mathbf{x}^\top \Sigma_{xx}^{-1}\mu_x - \frac{1}{2}\mu_x^\top \Sigma_{xx}^{-1}\mu_x \end{aligned} \quad (42)$$

- Using (24) in concert with (40) and (41), the middle term in α_2 can be written

$$\begin{aligned} \mathbf{x}^\top \Sigma_{xx}^{-1}\mu_x &= \mathbf{x}^\top (\Lambda_{xx} - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx}) \mu_x \\ &= \mathbf{x}^\top (\Lambda_{xx}\mu_x - \Lambda_{xy}\Lambda_{yy}^{-1}\Lambda_{yx}\mu_x) \\ &= \mathbf{x}^\top (\eta_x - \Lambda_{xy}\mu_y - \Lambda_{xy}\Lambda_{yy}^{-1}(\eta_y - \Lambda_{yy}\mu_y)) \\ &= \mathbf{x}^\top (\eta_x - \Lambda_{xy}\mu_y - \Lambda_{xy}\Lambda_{yy}^{-1}\eta_y + \Lambda_{xy}\mu_y) \\ &= \mathbf{x}^\top (\eta_x - \Lambda_{xy}\Lambda_{yy}^{-1}\eta_y) \\ &= \mathbf{x}^\top \bar{\eta}, \end{aligned} \quad (44)$$

where

$$\bar{\eta} = \eta_x - \Lambda_{xy}\Lambda_{yy}^{-1}\eta_y. \quad (45)$$

- ▶ The last term in α_2 can be written as

$$-\frac{1}{2}\boldsymbol{\mu}_x^\top \boldsymbol{\Sigma}_{xx}^{-1} \boldsymbol{\mu}_x \quad (46)$$

$$\begin{aligned} &= -\frac{1}{2}\boldsymbol{\mu}_x^\top \boldsymbol{\Sigma}_{xx}^{-1} \boldsymbol{\Sigma}_{xx} \boldsymbol{\Sigma}_{xx}^{-1} \boldsymbol{\mu}_x \\ &= -\frac{1}{2}(\boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\eta}_y)^\top (\boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx})^{-1} (\boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\eta}_y) \\ &= -\frac{1}{2}\bar{\boldsymbol{\eta}}^\top \bar{\boldsymbol{\Lambda}} \bar{\boldsymbol{\eta}}, \end{aligned} \quad (47)$$

where (45) has been used, and

$$\bar{\boldsymbol{\Lambda}} = \boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}. \quad (48)$$

- ▶ Thus, the exponent α_2 can be written as

$$\alpha_2 = -\frac{1}{2}\mathbf{x}^\top \boldsymbol{\Sigma}_{xx}^{-1} \mathbf{x} + \mathbf{x}^\top \bar{\boldsymbol{\eta}} - \frac{1}{2}\bar{\boldsymbol{\eta}}^\top \bar{\boldsymbol{\Lambda}} \bar{\boldsymbol{\eta}}. \quad (49)$$

- ▶ It follows that

$$p(\mathbf{x}) = \frac{\exp(-\frac{1}{2}\bar{\boldsymbol{\eta}}^\top \bar{\boldsymbol{\Lambda}} \bar{\boldsymbol{\eta}})}{\sqrt{(2\pi)^{n_x} \det \bar{\boldsymbol{\Lambda}}}} \exp\left(-\frac{1}{2}\mathbf{x}^\top \bar{\boldsymbol{\Lambda}} \mathbf{x} + \mathbf{x}^\top \bar{\boldsymbol{\eta}}\right) \quad (50)$$

$$= \mathcal{N}^{-1}(\bar{\boldsymbol{\eta}}, \bar{\boldsymbol{\Lambda}}), \quad (51)$$

where

$$\bar{\boldsymbol{\eta}} = \boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\eta}_y, \quad (45)$$

$$\bar{\boldsymbol{\Lambda}} = \boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}. \quad (48)$$

Conditioning Theorem

► Consider

$$\mathbf{z} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \quad \mathbf{z} = \begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}, \quad \boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xy} \\ \boldsymbol{\Sigma}_{yx} & \boldsymbol{\Sigma}_{yy} \end{bmatrix}, \quad \boldsymbol{\Lambda} = \boldsymbol{\Sigma}^{-1} = \begin{bmatrix} \boldsymbol{\Lambda}_{xx} & \boldsymbol{\Lambda}_{xy} \\ \boldsymbol{\Lambda}_{yx} & \boldsymbol{\Lambda}_{yy} \end{bmatrix}. \quad (18)$$

Theorem (4)

Let $\mathbf{z} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ and be partitioned as in (18). Then, the conditional pdf $p(\mathbf{x}|\mathbf{y})$ is given by

$$\begin{aligned} p(\mathbf{x}|\mathbf{y}) &= \mathcal{N}(\boldsymbol{\mu}_{x|y}, \boldsymbol{\Sigma}_{x|y}), \\ &= \mathcal{N}^{-1}(\bar{\boldsymbol{\eta}}_{x|y}, \bar{\boldsymbol{\Lambda}}_{x|y}), \end{aligned}$$

where

$$\boldsymbol{\mu}_{x|y} = \boldsymbol{\mu}_x + \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} (\mathbf{y} - \boldsymbol{\mu}_y), \quad \boldsymbol{\Sigma}_{x|y} = \boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx},$$

and

$$\bar{\boldsymbol{\eta}}_{x|y} = \boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \mathbf{y}, \quad \bar{\boldsymbol{\Lambda}}_{x|y} = \boldsymbol{\Lambda}_{xx}.$$

Proof.

- Recall that the relationship between $p(\mathbf{x}, \mathbf{y})$, $p(\mathbf{x}|\mathbf{y})$, and $p(\mathbf{y})$ is given by

$$p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{x}, \mathbf{y})}{p(\mathbf{y})}, \quad (52)$$

- The densities $p(\mathbf{y})$ and $p(\mathbf{x}, \mathbf{y})$ are

$$p(\mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}_y, \boldsymbol{\Sigma}_{yy}) = \frac{1}{\sqrt{(2\pi)^{n_y} \det \boldsymbol{\Sigma}_{yy}}} \exp \left(-\frac{1}{2} (\mathbf{y} - \boldsymbol{\mu}_y)^\top \boldsymbol{\Sigma}_{yy}^{-1} (\mathbf{y} - \boldsymbol{\mu}_y) \right), \quad (53)$$

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^{(n_x+n_y)} \det \boldsymbol{\Sigma}}} \exp(\alpha), \quad (54)$$

and the argument of the exponent, α , is

$$\alpha = -\frac{1}{2} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right)^\top \begin{bmatrix} \boldsymbol{\Lambda}_{xx} & \boldsymbol{\Lambda}_{xy} \\ \boldsymbol{\Lambda}_{yx} & \boldsymbol{\Lambda}_{yy} \end{bmatrix} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right). \quad (55)$$

- ▶ Using the Schur complement of Λ_{xx} in Λ , Theorem 2, enables α to be written as

$$\alpha = -\frac{1}{2} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right)^\top \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \Lambda_{yx} \Lambda_{xx}^{-1} & \mathbf{1} \end{bmatrix} \begin{bmatrix} \Lambda_{xx} & \mathbf{0} \\ \mathbf{0} & \Lambda_{yy} - \Lambda_{yx} \Lambda_{xx}^{-1} \Lambda_{xy} \end{bmatrix} \times \begin{bmatrix} \mathbf{1} & \Lambda_{xx}^{-1} \Lambda_{xy} \\ \mathbf{0} & \mathbf{1} \end{bmatrix} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix} - \begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix} \right). \quad (56)$$

- ▶ Next, expand α , and write α as the sum of two quadratic terms:

- ▶ a quadratic term in $(\mathbf{x} - \boldsymbol{\mu}_x) + \Lambda_{xx}^{-1} \Lambda_{xy}(\mathbf{y} - \boldsymbol{\mu}_y)$, called α_3 , and
- ▶ a quadratic term in $\mathbf{y} - \boldsymbol{\mu}_y$, called α_4 .

- ▶ It follows that $\alpha = \alpha_3 + \alpha_4$ where

$$\begin{aligned} \alpha_3 &= -\frac{1}{2} \left((\mathbf{x} - \boldsymbol{\mu}_x) + \Lambda_{xx}^{-1} \Lambda_{xy}(\mathbf{y} - \boldsymbol{\mu}_y) \right)^\top \Lambda_{xx} \left((\mathbf{x} - \boldsymbol{\mu}_x) + \Lambda_{xx}^{-1} \Lambda_{xy}(\mathbf{y} - \boldsymbol{\mu}_y) \right), \\ &= -\frac{1}{2} \left(\mathbf{x} - (\boldsymbol{\mu}_x - \Lambda_{xx}^{-1} \Lambda_{xy}(\mathbf{y} - \boldsymbol{\mu}_y)) \right)^\top \Lambda_{xx} \left(\mathbf{x} - (\boldsymbol{\mu}_x - \Lambda_{xx}^{-1} \Lambda_{xy}(\mathbf{y} - \boldsymbol{\mu}_y)) \right) \end{aligned} \quad (57)$$

$$\begin{aligned} \alpha_4 &= -\frac{1}{2} (\mathbf{y} - \boldsymbol{\mu}_y)^\top (\Lambda_{yy} - \Lambda_{yx} \Lambda_{xx}^{-1} \Lambda_{xy}) (\mathbf{y} - \boldsymbol{\mu}_y) \\ &= -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \Sigma_{yy}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x), \end{aligned} \quad (58)$$

and (26) has been used.

- ▶ Next, focus on writing α_3 as

$$\alpha_3 = -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_{x|y})^\top \boldsymbol{\Sigma}_{x|y}^{-1} (\mathbf{x} - \boldsymbol{\mu}_{x|y}) \quad (59)$$

where

$$\boldsymbol{\mu}_{x|y} = \boldsymbol{\mu}_x + \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} (\mathbf{y} - \boldsymbol{\mu}_y), \quad (60)$$

$$\boldsymbol{\Sigma}_{x|y} = \boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx}. \quad (61)$$

- ▶ From (20),

$$\boldsymbol{\Sigma}_{x|y}^{-1} = \boldsymbol{\Lambda}_{xx} = (\boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx})^{-1}. \quad (62)$$

- ▶ To write $\boldsymbol{\mu}_{x|y}$ in (60) in terms of $\boldsymbol{\Sigma}_{xy}$ and $\boldsymbol{\Sigma}_{yy}$, consider the following.

- ▶ Using (19), $\boldsymbol{\Lambda}_{xx}$ and $\boldsymbol{\Lambda}_{xy}$ are

$$\boldsymbol{\Lambda}_{xx} = (\boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx})^{-1}, \quad (63)$$

$$\boldsymbol{\Lambda}_{xy} = -(\boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx})^{-1} \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1}. \quad (64)$$

- ▶ As such, the matrix product $-\boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xy}$ can be written as

$$\begin{aligned} -\boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xy} &= (\boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx}) (\boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx})^{-1} \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \\ &= \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1}. \end{aligned} \quad (65)$$

- ▶ Thus

$$\boldsymbol{\mu}_{x|y} = \boldsymbol{\mu}_x - \boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xy} (\mathbf{y} - \boldsymbol{\mu}_y) = \boldsymbol{\mu}_x + \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} (\mathbf{y} - \boldsymbol{\mu}_y). \quad (66)$$

- ▶ It follows that α_3 can be written as in (59).

- The joint pdf $p(\mathbf{x}, \mathbf{y})$ can now be written as

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^{(n_x+n_y)} \det \boldsymbol{\Sigma}}} \exp(\alpha_3) \exp(\alpha_4) \quad (67)$$

- It follows that

$$\begin{aligned} p(\mathbf{x}|\mathbf{y}) &= p(\mathbf{x}, \mathbf{y}) \left(\frac{1}{p(\mathbf{y})} \right) \\ &= \frac{1}{\sqrt{(2\pi)^{(n_x+n_y)} \det \boldsymbol{\Sigma}}} \exp(\alpha_3) \exp(\alpha_4) \left(\sqrt{(2\pi)^{n_y} \det \boldsymbol{\Sigma}_{yy}} \exp(-\alpha_4) \right) \\ &= \frac{\sqrt{\det \boldsymbol{\Sigma}_{yy}}}{\sqrt{(2\pi)^{n_x} \det \boldsymbol{\Sigma}}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_{x|y})^\top \boldsymbol{\Sigma}_{x|y}^{-1} (\mathbf{x} - \boldsymbol{\mu}_{x|y})\right). \end{aligned} \quad (68)$$

- ▶ Focusing on the normalization constant, using (17), first note that ,

$$\det \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix} = \det \Sigma_{yy} \det(\Sigma_{xx} - \Sigma_{xy} \Sigma_{yy}^{-1} \Sigma_{yx}). \quad (69)$$

- ▶ Then, the normalization constant can be simplified as

$$\begin{aligned} \frac{\sqrt{\det \Sigma_{yy}}}{\sqrt{(2\pi)^{n_x} \det \Sigma}} &= \frac{\sqrt{\det \Sigma_{yy}}}{\sqrt{(2\pi)^{n_x} \det \Sigma_{yy} \det(\Sigma_{xx} - \Sigma_{xy} \Sigma_{yy}^{-1} \Sigma_{yx})}} \\ &= \frac{1}{\sqrt{(2\pi)^{n_x} \det(\Sigma_{xx} - \Sigma_{xy} \Sigma_{yy}^{-1} \Sigma_{yx})}}. \end{aligned} \quad (70)$$

- ▶ Therefore,

$$p(\mathbf{x}|\mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}_{x|y}, \Sigma_{x|y}) = \frac{1}{\sqrt{(2\pi)^{n_x} \det \Sigma_{x|y}}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_{x|y})^\top \Sigma_{x|y}^{-1} (\mathbf{x} - \boldsymbol{\mu}_{x|y})\right), \quad (71)$$

where

$$\boldsymbol{\mu}_{x|y} = \boldsymbol{\mu}_x + \Sigma_{xy} \Sigma_{yy}^{-1} (\mathbf{y} - \boldsymbol{\mu}_y), \quad (60)$$

$$\Sigma_{x|y} = \Lambda_{xx}^{-1} = \Sigma_{xx} - \Sigma_{xy} \Sigma_{yy}^{-1} \Sigma_{yx}. \quad (61)$$

- ▶ To derive the information form, first recall that

$$\begin{bmatrix} \eta_x \\ \eta_y \end{bmatrix} = \begin{bmatrix} \Lambda_{xx} & \Lambda_{xy} \\ \Lambda_{yx} & \Lambda_{yy} \end{bmatrix} \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix} = \begin{bmatrix} \Lambda_{xx}\mu_x + \Lambda_{xy}\mu_y \\ \Lambda_{yx}\mu_x + \Lambda_{yy}\mu_y \end{bmatrix},$$

or

$$\Lambda_{xx}\mu_x = \eta_x - \Lambda_{xy}\mu_y \quad (40)$$

$$\Lambda_{yx}\mu_x = \eta_y - \Lambda_{yy}\mu_y. \quad (41)$$

- ▶ Expanding the exponent, α_3 , results in

$$\begin{aligned} \alpha_3 &= -\frac{1}{2} (\mathbf{x} - \mu_{x|y})^\top \Sigma_{x|y}^{-1} (\mathbf{x} - \mu_{x|y}) \\ &= -\frac{1}{2} \mathbf{x}^\top \Sigma_{x|y}^{-1} \mathbf{x} + \mathbf{x}^\top \Sigma_{x|y}^{-1} \mu_{x|y} - \frac{1}{2} \mu_{x|y}^\top \Sigma_{x|y}^{-1} \mu_{x|y}. \end{aligned} \quad (72)$$

- ▶ Using (62), (66), and (40), the middle term in α_3 can be written as

$$\mathbf{x}^\top \Sigma_{x|y}^{-1} \mu_{x|y} = \mathbf{x}^\top \Lambda_{xx} (\mu_x - \Lambda_{xx}^{-1} \Lambda_{xy} (\mathbf{y} - \mu_y)) \quad (73)$$

$$\begin{aligned} &= \mathbf{x}^\top (\Lambda_{xx}\mu_x - \Lambda_{xy}(\mathbf{y} - \mu_y)) \\ &= \mathbf{x}^\top (\eta_x - \Lambda_{xy}\mu_y - \Lambda_{xy}\mathbf{y} + \Lambda_{xy}\mu_y) \\ &= \mathbf{x}^\top (\eta_x - \Lambda_{xy}\mathbf{y}) \\ &= \mathbf{x}^\top \bar{\eta}_{x|y}, \end{aligned} \quad (74)$$

where

$$\bar{\eta}_{x|y} = \eta_x - \Lambda_{xy}\mathbf{y}. \quad (75)$$

- The last term in α_3 can be written as

$$- \frac{1}{2} \boldsymbol{\mu}_{x|y}^\top \boldsymbol{\Sigma}_{x|y}^{-1} \boldsymbol{\mu}_{x|y} \quad (76)$$

$$\begin{aligned} &= -\frac{1}{2} \boldsymbol{\mu}_{x|y}^\top \boldsymbol{\Lambda}_{xx} \boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xx} \boldsymbol{\mu}_{x|y} \\ &= -\frac{1}{2} \left(\boldsymbol{\mu}_x - \boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xy} (\mathbf{y} - \boldsymbol{\mu}_y) \right)^\top \boldsymbol{\Lambda}_{xx} \boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xx} \left(\boldsymbol{\mu}_x - \boldsymbol{\Lambda}_{xx}^{-1} \boldsymbol{\Lambda}_{xy} (\mathbf{y} - \boldsymbol{\mu}_y) \right) \\ &= -\bar{\boldsymbol{\eta}}_{x|y}^\top \bar{\boldsymbol{\Lambda}}_{x|y}^{-1} \bar{\boldsymbol{\eta}}_{x|y} \end{aligned} \quad (77)$$

where

$$\bar{\boldsymbol{\Lambda}}_{x|y} = \boldsymbol{\Lambda}_{xx}. \quad (78)$$

- Thus, the exponent α_3 can be written as

$$\begin{aligned} \alpha_3 &= -\frac{1}{2} \mathbf{x}^\top \boldsymbol{\Sigma}_{x|y}^{-1} \mathbf{x} + \mathbf{x}^\top \boldsymbol{\Sigma}_{x|y}^{-1} \boldsymbol{\mu}_{x|y} - \frac{1}{2} \boldsymbol{\mu}_{x|y}^\top \boldsymbol{\Sigma}_{x|y}^{-1} \boldsymbol{\mu}_{x|y} \\ &= -\frac{1}{2} \mathbf{x}^\top \bar{\boldsymbol{\Lambda}}_{x|y} \mathbf{x} + \mathbf{x}^\top \bar{\boldsymbol{\eta}}_{x|y} - \bar{\boldsymbol{\eta}}_{x|y}^\top \bar{\boldsymbol{\Lambda}}_{x|y}^{-1} \bar{\boldsymbol{\eta}}_{x|y}. \end{aligned} \quad (79)$$

- It follows that

$$\begin{aligned} p(\mathbf{x}|\mathbf{y}) &= \frac{\exp(-\bar{\boldsymbol{\eta}}_{x|y}^\top \bar{\boldsymbol{\Lambda}}_{x|y}^{-1} \bar{\boldsymbol{\eta}}_{x|y})}{\sqrt{(2\pi)^{n_x} \det \bar{\boldsymbol{\Lambda}}_{x|y}}} \exp\left(-\frac{1}{2} \mathbf{x}^\top \bar{\boldsymbol{\Lambda}}_{x|y} \mathbf{x} + \mathbf{x}^\top \bar{\boldsymbol{\eta}}_{x|y}\right) \\ &= \mathcal{N}^{-1}(\bar{\boldsymbol{\eta}}_{x|y}, \bar{\boldsymbol{\Lambda}}_{x|y}), \end{aligned}$$

where

$$\bar{\boldsymbol{\eta}}_{x|y} = \boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \mathbf{y}, \quad (75)$$

$$\bar{\boldsymbol{\Lambda}}_{x|y} = \boldsymbol{\Lambda}_{xx}. \quad (78)$$

Summary

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N} \left(\begin{bmatrix} \boldsymbol{\mu}_x \\ \boldsymbol{\mu}_y \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xy} \\ \boldsymbol{\Sigma}_{yx} & \boldsymbol{\Sigma}_{yy} \end{bmatrix} \right) = \mathcal{N}^{-1} \left(\begin{bmatrix} \boldsymbol{\eta}_x \\ \boldsymbol{\eta}_y \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Lambda}_{xx} & \boldsymbol{\Lambda}_{xy} \\ \boldsymbol{\Lambda}_{yx} & \boldsymbol{\Lambda}_{yy} \end{bmatrix} \right)$$

	Marginalization $p(\mathbf{x}) = \int_{-\infty}^{\infty} p(\mathbf{x}, \mathbf{y}) d\mathbf{y}$	Conditioning $p(\mathbf{x}, \mathbf{y}) = p(\mathbf{x} \mathbf{y})p(\mathbf{y})$
Cov. Form	$\boldsymbol{\mu} = \boldsymbol{\mu}_x$ $\boldsymbol{\Sigma} = \boldsymbol{\Sigma}_{xx}$	$\boldsymbol{\mu}_{x y} = \boldsymbol{\mu}_x + \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} (\mathbf{y} - \boldsymbol{\mu}_y)$ $\boldsymbol{\Sigma}_{x y} = \boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{yy}^{-1} \boldsymbol{\Sigma}_{yx}$
Info. Form	$\bar{\boldsymbol{\eta}} = \boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\eta}_y$ $\bar{\boldsymbol{\Lambda}} = \boldsymbol{\Lambda}_{xx} - \boldsymbol{\Lambda}_{xy} \boldsymbol{\Lambda}_{yy}^{-1} \boldsymbol{\Lambda}_{yx}$	$\bar{\boldsymbol{\eta}}_{x y} = \boldsymbol{\eta}_x - \boldsymbol{\Lambda}_{xy} \mathbf{y}$ $\bar{\boldsymbol{\Lambda}}_{x y} = \boldsymbol{\Lambda}_{xx}$

(Eustice *et al.* 2006, IEEE Trans. Robotics)

Joint Density with Affine Transformation (Lemma A.1 in Särkkä)

Theorem (5)

Consider $\mathbf{x} \in \mathbb{R}^{n_x}$ and $\mathbf{y} \in \mathbb{R}^{n_y}$ where $\mathbf{y} = \mathbf{M}\mathbf{x} + \mathbf{b}$, $\mathbf{M} \in \mathbb{R}^{n_y \times n_x}$ is constant, and $\mathbf{b} \in \mathbb{R}^{n_y}$ is constant. Let

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_x, \boldsymbol{\Sigma}_{xx}), \quad p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}; \mathbf{M}\mathbf{x} + \mathbf{b}, \boldsymbol{\Sigma}_{y|x}).$$

Then, the joint density of $\mathbf{x}$ and $\mathbf{y}$, $p(\mathbf{x}, \mathbf{y})$, is

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}; \begin{bmatrix} \boldsymbol{\mu}_x \\ \mathbf{M}\boldsymbol{\mu}_x + \mathbf{b} \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top \\ \mathbf{M}\boldsymbol{\Sigma}_{xx} & \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top + \boldsymbol{\Sigma}_{y|x} \end{bmatrix} \right).$$

Proof.

► Recall that

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{y}|\mathbf{x})p(\mathbf{x}). \quad (80)$$

► The task at hand is to compute

$$p(\mathbf{x}, \mathbf{y}) \text{ given } p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}; \mathbf{M}\mathbf{x} + \mathbf{b}, \Sigma_{y|x}), \quad p(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_x, \Sigma_{xx}). \quad (81)$$

► By direct computation,

$$\begin{aligned} p(\mathbf{x}, \mathbf{y}) &= p(\mathbf{y}|\mathbf{x})p(\mathbf{x}) \\ &= \frac{1}{\sqrt{(2\pi)^{n_y} \det \Sigma_{y|x}}} \exp \left(-\frac{1}{2} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b}))^\top \Sigma_{y|x}^{-1} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b})) \right) \\ &\quad \times \frac{1}{\sqrt{(2\pi)^{n_x} \det \Sigma_{xx}}} \exp \left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \Sigma_{xx}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x) \right) \\ &= \frac{1}{\sqrt{(2\pi)^{n_x+n_y} \det \Sigma_{y|x} \det \Sigma_{xx}}} \exp(\alpha), \end{aligned} \quad (82)$$

where

$$\alpha = -\frac{1}{2} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b}))^\top \Sigma_{y|x}^{-1} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b})) - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \Sigma_{xx}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x). \quad (83)$$

► Write $\alpha = \alpha_5 + \alpha_6$ where

$$\alpha_5 = -\frac{1}{2} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b}))^\top \Sigma_{y|x}^{-1} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b})), \quad (84)$$

$$\alpha_6 = -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \Sigma_{xx}^{-1} (\mathbf{x} - \boldsymbol{\mu}_x). \quad (85)$$

► Add and subtract $\mathbf{M}\boldsymbol{\mu}_x$ in the first term of α_5 and rearrange to yield

$$\begin{aligned} \alpha_5 &= -\frac{1}{2} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b}) + \mathbf{M}\boldsymbol{\mu}_x - \mathbf{M}\boldsymbol{\mu}_x)^\top \Sigma_{y|x}^{-1} (\mathbf{y} - (\mathbf{M}\mathbf{x} + \mathbf{b}) + \mathbf{M}\boldsymbol{\mu}_x - \mathbf{M}\boldsymbol{\mu}_x) \\ &= -\frac{1}{2} (\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) - \mathbf{M}(\mathbf{x} - \boldsymbol{\mu}_x))^\top \Sigma_{y|x}^{-1} (\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) - \mathbf{M}(\mathbf{x} - \boldsymbol{\mu}_x)) \\ &= -\frac{1}{2} (\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}))^\top \Sigma_{y|x}^{-1} (\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b})) \\ &\quad + (\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}))^\top \Sigma_{y|x}^{-1} \mathbf{M}(\mathbf{x} - \boldsymbol{\mu}_x) \\ &\quad - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_x)^\top \mathbf{M}^\top \Sigma_{y|x}^{-1} \mathbf{M}(\mathbf{x} - \boldsymbol{\mu}_x) \\ &= -\frac{1}{2} \begin{bmatrix} \mathbf{x} - \boldsymbol{\mu}_x \\ \mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) \end{bmatrix}^\top \begin{bmatrix} \mathbf{M}^\top \Sigma_{y|x}^{-1} \mathbf{M} & -\mathbf{M}^\top \Sigma_{y|x}^{-1} \\ -\Sigma_{y|x}^{-1} \mathbf{M} & \Sigma_{y|x}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{x} - \boldsymbol{\mu}_x \\ \mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) \end{bmatrix}. \end{aligned} \quad (86)$$

- It follows that α can be written as

$$\alpha = -\frac{1}{2} \begin{bmatrix} \mathbf{x} - \boldsymbol{\mu}_x \\ \mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) \end{bmatrix}^T \begin{bmatrix} \boldsymbol{\Sigma}_{xx}^{-1} + \mathbf{M}^T \boldsymbol{\Sigma}_{y|x}^{-1} \mathbf{M} & -\mathbf{M}^T \boldsymbol{\Sigma}_{y|x}^{-1} \\ -\boldsymbol{\Sigma}_{y|x}^{-1} \mathbf{M} & \boldsymbol{\Sigma}_{y|x}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{x} - \boldsymbol{\mu}_x \\ \mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) \end{bmatrix} \quad (87)$$

- Next, partition the information matrix as

$$\mathbf{X}^{-1} = \begin{bmatrix} \boldsymbol{\Sigma}_{xx}^{-1} + \mathbf{M}^T \boldsymbol{\Sigma}_{y|x}^{-1} \mathbf{M} & -\mathbf{M}^T \boldsymbol{\Sigma}_{y|x}^{-1} \\ -\boldsymbol{\Sigma}_{y|x}^{-1} \mathbf{M} & \boldsymbol{\Sigma}_{y|x}^{-1} \end{bmatrix} \quad (88)$$

$$= \begin{bmatrix} \mathbf{1} & -\mathbf{M}^T \\ \mathbf{0} & \mathbf{1} \end{bmatrix} \begin{bmatrix} \boldsymbol{\Sigma}_{xx}^{-1} & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Sigma}_{y|x}^{-1} \end{bmatrix} \begin{bmatrix} \mathbf{1} & -\mathbf{M}^T \\ -\mathbf{M} & \mathbf{1} \end{bmatrix}. \quad (89)$$

- Using the fact that if $\mathbf{X} = \mathbf{PQR}$, then $\mathbf{X}^{-1} = \mathbf{R}^{-1}\mathbf{Q}^{-1}\mathbf{P}^{-1}$, it follows that the covariance matrix is

$$\begin{aligned} \mathbf{X} &= \begin{bmatrix} \mathbf{1} & \mathbf{0} \\ \mathbf{M} & \mathbf{1} \end{bmatrix} \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Sigma}_{y|x} \end{bmatrix} \begin{bmatrix} \mathbf{1} & \mathbf{M}^T \\ \mathbf{0} & \mathbf{1} \end{bmatrix} \\ &= \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xx} \mathbf{M}^T \\ \mathbf{M} \boldsymbol{\Sigma}_{xx} & \mathbf{M} \boldsymbol{\Sigma}_{xx} \mathbf{M}^T + \boldsymbol{\Sigma}_{y|x} \end{bmatrix}. \end{aligned} \quad (90)$$

► To conclude,

$$\begin{aligned} p(\mathbf{x}, \mathbf{y}) &= p(\mathbf{y}|\mathbf{x})p(\mathbf{x}) \\ &= \frac{1}{\sqrt{(2\pi)^{n_x+n_y} \det \Sigma_{y|x} \det \Sigma_{xx}}} \exp(\alpha), \end{aligned} \quad (91)$$

where

$$\alpha = -\frac{1}{2} \left[\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) \right]^\top \begin{bmatrix} \Sigma_{xx} & \Sigma_{xx}\mathbf{M}^\top \\ \mathbf{M}\Sigma_{xx} & \mathbf{M}\Sigma_{xx}\mathbf{M}^\top + \Sigma_{y|x} \end{bmatrix}^{-1} \left[\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}) \right]. \quad (92)$$

Marginal and Cond. Densities with Affine Trans.

Theorem (6)

Consider $\mathbf{x} \in \mathbb{R}^{n_x}$ and $\mathbf{y} \in \mathbb{R}^{n_y}$ where $\mathbf{y} = \mathbf{M}\mathbf{x} + \mathbf{b}$, $\mathbf{M} \in \mathbb{R}^{n_y \times n_x}$ is constant, and $\mathbf{b} \in \mathbb{R}^{n_y}$ is constant. Let

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_x, \boldsymbol{\Sigma}_{xx}), \quad p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}; \mathbf{M}\mathbf{x} + \mathbf{b}, \boldsymbol{\Sigma}_{y|x}).$$

Then, the marginal density of $\mathbf{y}$, $p(\mathbf{y})$, is

$$p(\mathbf{y}) = \mathcal{N}(\mathbf{y}; \boldsymbol{\mu}_y, \boldsymbol{\Sigma}_y),$$

where

$$\boldsymbol{\mu}_y = \mathbf{M}\boldsymbol{\mu}_x + \mathbf{b},$$

$$\boldsymbol{\Sigma}_y = \boldsymbol{\Sigma}_{y|x} + \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top,$$

and the conditional density of $\mathbf{x}$ given $\mathbf{y}$, $p(\mathbf{x}|\mathbf{y})$, is

$$p(\mathbf{x}|\mathbf{y}) = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_{x|y}, \boldsymbol{\Sigma}_{x|y}),$$

where

$$\boldsymbol{\mu}_{x|y} = \boldsymbol{\mu}_x + \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top\boldsymbol{\Sigma}_y^{-1}(\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b})),$$

$$\boldsymbol{\Sigma}_{x|y} = \boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top\boldsymbol{\Sigma}_y^{-1}\mathbf{M}\boldsymbol{\Sigma}_{xx}.$$

Proof.

- Recall that

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{y}|\mathbf{x})p(\mathbf{x}). \quad (93)$$

- Using Theorem 5 (Joint Density with Affine Transformation Theorem),

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N} \left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}; \begin{bmatrix} \boldsymbol{\mu}_x \\ \mathbf{M}\boldsymbol{\mu}_x + \mathbf{b} \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top \\ \mathbf{M}\boldsymbol{\Sigma}_{xx} & \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top + \boldsymbol{\Sigma}_{y|x} \end{bmatrix} \right). \quad (94)$$

- Using Theorem 3 (Marginalization Theorem),

$$p(\mathbf{y}) = \mathcal{N}(\mathbf{y}; \boldsymbol{\mu}_y, \boldsymbol{\Sigma}_y) \quad (95)$$

where

$$\begin{aligned} \boldsymbol{\mu}_y &= \mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}, \\ \boldsymbol{\Sigma}_y &= \boldsymbol{\Sigma}_{y|x} + \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top. \end{aligned}$$

► Next, recall that

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{x}|\mathbf{y})p(\mathbf{y}), \quad \text{or} \quad p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{x}, \mathbf{y})}{p(\mathbf{y})}. \quad (96)$$

► Recall that

$$p(\mathbf{y}) = \mathcal{N}(\mathbf{y}; \mathbf{M}\boldsymbol{\mu}_x + \mathbf{b}, \boldsymbol{\Sigma}_{y|x} + \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top), \quad (97)$$

$$p(\mathbf{x}, \mathbf{y}) = \mathcal{N}\left(\begin{bmatrix} \mathbf{x} \\ \mathbf{y} \end{bmatrix}; \begin{bmatrix} \boldsymbol{\mu}_x \\ \mathbf{M}\boldsymbol{\mu}_x + \mathbf{b} \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{xx} & \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top \\ \mathbf{M}\boldsymbol{\Sigma}_{xx} & \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top + \boldsymbol{\Sigma}_{y|x} \end{bmatrix}\right). \quad (98)$$

► Using Theorem 4 (Conditioning Theorem),

$$p(\mathbf{x}|\mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}_{x|y}, \boldsymbol{\Sigma}_{x|y}),$$

where

$$\begin{aligned} \boldsymbol{\mu}_{x|y} &= \boldsymbol{\mu}_x + \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top(\boldsymbol{\Sigma}_{y|x} + \mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top)^{-1}(\mathbf{y} - (\mathbf{M}\boldsymbol{\mu}_x + \mathbf{b})), \\ \boldsymbol{\Sigma}_{x|y} &= \boldsymbol{\Sigma}_{xx} - \boldsymbol{\Sigma}_{xx}\mathbf{M}^\top(\mathbf{M}\boldsymbol{\Sigma}_{xx}\mathbf{M}^\top + \boldsymbol{\Sigma}_{y|x})^{-1}\mathbf{M}\boldsymbol{\Sigma}_{xx}. \end{aligned}$$

Bayes' Rule

- ▶ The joint pdf $p(\mathbf{x}, \mathbf{y})$ can be written as

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{x}|\mathbf{y})p(\mathbf{y}), \quad \text{or} \quad (99)$$

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{y}|\mathbf{x})p(\mathbf{x}). \quad (100)$$

- ▶ Setting (99) equal to (100) and rearranging leads to Bayes' rule,

$$p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{x})p(\mathbf{x})}{p(\mathbf{y})}. \quad (101)$$

- ▶ Write $p(\mathbf{y})$ as

$$\begin{aligned} p(\mathbf{y}) &= p(\mathbf{y}) \underbrace{\int_a^b p(\mathbf{x}|\mathbf{y}) \, d\mathbf{x}}_{=1} = \int_a^b p(\mathbf{x}|\mathbf{y})p(\mathbf{y}) \, d\mathbf{x} = \int_a^b p(\mathbf{x}, \mathbf{y}) \, d\mathbf{x} \\ &= \int_a^b p(\mathbf{y}|\mathbf{x})p(\mathbf{x}) \, d\mathbf{x}. \end{aligned}$$

- ▶ It follows that Bayes' rule can then be written as

$$p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{x})p(\mathbf{x})}{\int_a^b p(\mathbf{y}|\mathbf{x})p(\mathbf{x}) \, d\mathbf{x}}.$$

- ▶ This applies to non-Gaussian, as well as Gaussian, distributions.

- ▶ In the Gaussian case, this is in fact exactly what was done in Theorem 6, the Marginal and Conditional Densities with Affine Transformation Theorem.
- ▶ Given $p(\mathbf{y}|\mathbf{x})$ and $p(\mathbf{x})$, construct the joint density $p(\mathbf{x}, \mathbf{y})$ using

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{y}|\mathbf{x})p(\mathbf{x}).$$

- ▶ Using Theorem 3, the Marginalization Theorem, marginalize out $\mathbf{x}$ from $p(\mathbf{x}, \mathbf{y})$ to get $p(\mathbf{y})$.
- ▶ Given the joint density $p(\mathbf{x}, \mathbf{y})$ and $p(\mathbf{y})$, each of which have just been computed, compute

$$p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{x}, \mathbf{y})}{p(\mathbf{y})}$$

using Theorem 4, the Conditioning Theorem.

References

Slides are based on [2–6].

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